

# Lecture 7: Model Selection and Regularization

## STATS 202: Data Mining and Analysis

Linh Tran

tranlm@stanford.edu



Department of Statistics  
Stanford University

July 17, 2023



- ▶ HW1 is graded.
  - ▶ Regrade requests are allowed up to 1 week from publication
- ▶ HW2 is due today
- ▶ Midterm is this Wednesday
  - ▶ In-person
  - ▶ Accommodations should be confirmed
  - ▶ No calculators necessary
- ▶ Anonymous course survey will be posted on Wednesday
  - ▶ Closes Tuesday afternoon
  - ▶ Up to 10 additional bonus points on exam (conditional on % participation)
- ▶ No section this Friday



- ▶ Subset selection
- ▶ Shrinkage methods
  - ▶ Ridge
  - ▶ LASSO
  - ▶ Elastic net



- ▶ In linear regression, adding predictors always decreases the training error or RSS.

$$RSS = (\mathbf{Y} - \mathbf{X}\beta)^\top (\mathbf{Y} - \mathbf{X}\beta) \quad (1)$$

- ▶ We can estimate  $\beta$  by minimizing the RSS.

$$\hat{\beta} = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top y \quad (2)$$

- ▶ However, adding predictors does not necessarily improve the test error.
- ▶ Selecting significant predictors is hard when  $n$  is not much larger than  $p$ .



- ▶ When our matrix is not of *full column rank* (e.g.  $n < p$ ), we have that  $(\mathbf{X}^\top \mathbf{X})^{-1}$  is not invertible.
- ▶ Consequently, there is no least squares solution:

$$\hat{\beta} = \underbrace{(\mathbf{X}^\top \mathbf{X})^{-1}}_{\text{Singular}} \mathbf{X}^\top y \quad (3)$$

- ▶ So, we must find a way around this.



$$\hat{\beta} = \underbrace{(\mathbf{X}^T \mathbf{X})^{-1}}_{\text{Singular}} \mathbf{X}^T y \quad (4)$$

Three common approaches for dealing with this:

## 1. Subset selection

- ▶ Select a subset  $k$  of the  $p$  predictors ( $k \leq p$ ).
- ▶ Use criteria to help select which subset  $k$  we want.



$$\hat{\beta} = \underbrace{(\mathbf{X}^T \mathbf{X})^{-1}}_{\text{Singular}} \mathbf{X}^T y \quad (4)$$

Three common approaches for dealing with this:

## 1. Subset selection

- ▶ Select a subset  $k$  of the  $p$  predictors ( $k \leq p$ ).
- ▶ Use criteria to help select which subset  $k$  we want.

## 2. Shrinkage methods

- ▶ Constrain the parameters we're estimating in some way



$$\hat{\beta} = \underbrace{(\mathbf{X}^T \mathbf{X})^{-1}}_{\text{Singular}} \mathbf{X}^T y \quad (4)$$

Three common approaches for dealing with this:

## 1. Subset selection

- ▶ Select a subset  $k$  of the  $p$  predictors ( $k \leq p$ ).
- ▶ Use criteria to help select which subset  $k$  we want.

## 2. Shrinkage methods

- ▶ Constrain the parameters we're estimating in some way

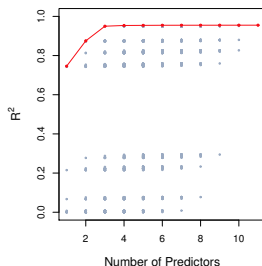
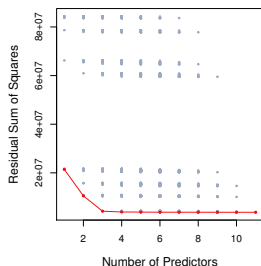
## 3. Dimension reduction

- ▶ Project all our predictors to a smaller dimension space
- ▶ Not covered in this class





- ▶ *Simple idea*: Compare all models with  $k$  predictors
- ▶ **Note**: There are  $\binom{p}{k} = p!/(k!(p-k)!)$  possible models
- ▶ Choose the model with the smallest RSS
  - ▶ Doing this for every possible  $k$ :



Note: As expected, the RSS and  $R^2$  improve with higher  $k$ .



## Two approaches:

1. Use a hold out set (e.g. validation or test set)
  - ▶ c.f. Cross-validation
2. Use *modified* metrics that account for the size of  $k$ , e.g.
  - ▶ Akaike Information Criterion (AIC)
  - ▶ Bayesian Information Criterion (BIC)
  - ▶ Adjusted  $R^2$



## Two approaches:

1. Use a hold out set (e.g. validation or test set)
  - ▶ c.f. Cross-validation
2. Use *modified* metrics that account for the size of  $k$ , e.g.
  - ▶ Akaike Information Criterion (AIC)
  - ▶ Bayesian Information Criterion (BIC)
  - ▶ Adjusted  $R^2$

How the modified metrics compare to using hold out sets

- ▶ Can be (much) less expensive to compute
- ▶ Motivated by asymptotic arguments and rely on model assumptions (e.g. normality of the errors)
- ▶ Equivalent concepts for other models (e.g. logistic regression)



Similar to Mallows's  $C_p$ :

$$C_p = \frac{1}{n}(RSS + 2k\hat{\sigma}^2) \quad (5)$$

- ▶ i.e. Adds the penalty  $2k\hat{\sigma}^2$  to the RSS
- ▶ Can be shown to be unbiased estimate of test set error

But, also normalizes for  $\hat{\sigma}^2$ :

$$AIC = \frac{1}{n\hat{\sigma}^2}(RSS + 2k\hat{\sigma}^2) = \frac{C_p}{\hat{\sigma}^2} \quad (6)$$

Since the two are proportional, (for least squares models) both are optimized at the same  $k$ .



Similar to Mallows's  $C_p$ , but derived from Bayesian POV:

$$BIC = \frac{1}{n}(RSS + \log(n)k\hat{\sigma}^2) \quad (7)$$

n.b.  $\log(n) > 2$  for  $n > 7$

- ▶ BIC will penalize more for large  $k$  (i.e. optimizes for smaller  $k$ )



**Recall:**

$$R^2 = 1 - \frac{RSS}{TSS} \quad (8)$$

The adjusted  $R^2$  penalizes for larger  $k$ :

$$R_{adj}^2 = 1 - \frac{RSS/(n - d - 1)}{TSS/(n - 1)} \quad (9)$$



**Recall:**

$$R^2 = 1 - \frac{RSS}{TSS} \quad (8)$$

The adjusted  $R^2$  penalizes for larger  $k$ :

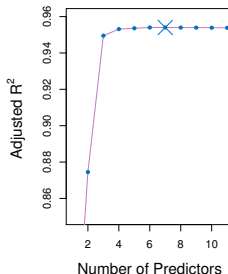
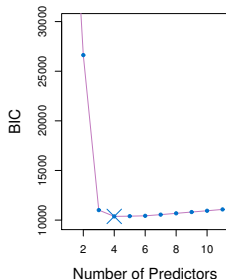
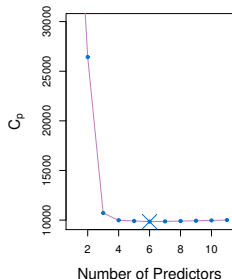
$$R_{adj}^2 = 1 - \frac{RSS/(n-d-1)}{TSS/(n-1)} \quad (9)$$

Maximizing  $R_{adj}^2$  is equivalent to minimizing  $1 - R_{adj}^2$ , i.e.:

$$\frac{RSS}{n-d-1} \quad (10)$$



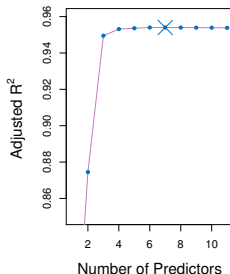
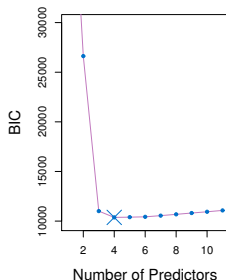
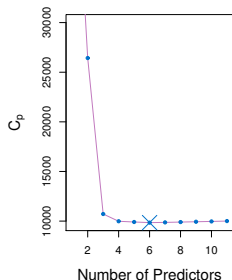
## Best subset selection for the Credit data set







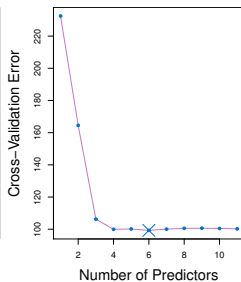
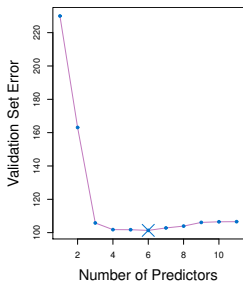
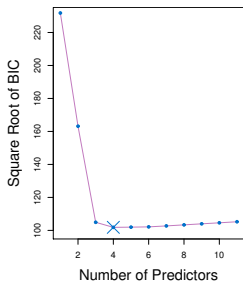
## Best subset selection for the Credit data set



n.b. The curve is pretty flat for  $k \geq 4$



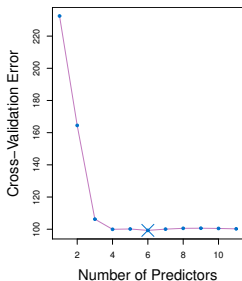
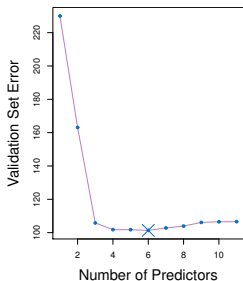
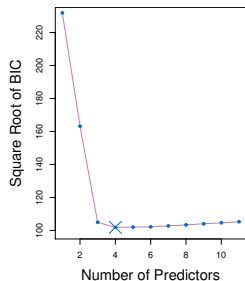
## BIC vs validation sets



n.b. The curves are also pretty flat for  $k \geq 4$ .



## BIC vs validation sets



n.b. The curves are also pretty flat for  $k \geq 4$ .

Can use the *one-standard-error rule*

- ▶ Choose the parsimonious model (i.e. lowest  $k$ ) such that the test error is within 1-SE of the lowest point



Best subset selection has 2 problems:

1. It is often very expensive computationally. We have to fit  $2^p$  different models!
2. If for a fixed  $k$ , there are too many possibilities, we increase our chances of overfitting
  - ▶ i.e. the model selected has high variance.



Best subset selection has 2 problems:

1. It is often very expensive computationally. We have to fit  $2^p$  different models!
2. If for a fixed  $k$ , there are too many possibilities, we increase our chances of overfitting
  - ▶ i.e. the model selected has high variance.

**One solution:** Restrict our search space for the best model

- ▶ This reduces the variance of the selected model at the expense of an increase in bias.



---

**Algorithm 6.2** *Forward stepwise selection*

---

1. Let  $\mathcal{M}_0$  denote the *null* model, which contains no predictors.
  2. For  $k = 0, \dots, p - 1$ :
    - (a) Consider all  $p - k$  models that augment the predictors in  $\mathcal{M}_k$  with one additional predictor.
    - (b) Choose the *best* among these  $p - k$  models, and call it  $\mathcal{M}_{k+1}$ . Here *best* is defined as having smallest RSS or highest  $R^2$ .
  3. Select a single best model from among  $\mathcal{M}_0, \dots, \mathcal{M}_p$  using cross-validated prediction error,  $C_p$  (AIC), BIC, or adjusted  $R^2$ .
-



# Variables	Best subset	Forward stepwise
One	<code>rating</code>	<code>rating</code>
Two	<code>rating, income</code>	<code>rating, income</code>
Three	<code>rating, income, student</code>	<code>rating, income, student</code>
Four	<code>cards, income</code> <code>student, limit</code>	<code>rating, income,</code> <code>student, limit</code>

**TABLE 6.1.** *The first four selected models for best subset selection and forward stepwise selection on the `Credit` data set. The first three models are identical but the fourth models differ.*



---

**Algorithm 6.3** *Backward stepwise selection*

---

1. Let  $\mathcal{M}_p$  denote the *full* model, which contains all  $p$  predictors.
  2. For  $k = p, p - 1, \dots, 1$ :
    - (a) Consider all  $k$  models that contain all but one of the predictors in  $\mathcal{M}_k$ , for a total of  $k - 1$  predictors.
    - (b) Choose the *best* among these  $k$  models, and call it  $\mathcal{M}_{k-1}$ . Here *best* is defined as having smallest RSS or highest  $R^2$ .
  3. Select a single best model from among  $\mathcal{M}_0, \dots, \mathcal{M}_p$  using cross-validated prediction error,  $C_p$  (AIC), BIC, or adjusted  $R^2$ .
-





- ▶ You cannot apply backward selection when  $p > n$ 
  - ▶ Though should still have a “reasonable” number of observations
- ▶ **Important:** they may not produce the same sequence of models.

Example:  $X_1, X_2 \stackrel{iid}{\sim} \mathcal{N}(0, \sigma^2)$

$$X_3 = X_1 + 3X_2 \quad (11)$$

$$Y = X_1 + 2X_2 + \epsilon \quad (12)$$

Regressing  $Y$  onto  $X_1, X_2, X_3$ :

- ▶ Forward:  $\{X_3\} \rightarrow \{X_3, X_2\} \rightarrow \{X_3, X_2, X_1\}$
- ▶ Backward:  $\{X_1, X_2, X_3\} \rightarrow \{X_1, X_2\} \rightarrow \{X_2\}$



- ▶ *Mixed stepwise selection*: Do forward selection, but at every step, remove any variables that are no longer “necessary”
  - ▶ e.g. using p-values
- ▶ *Forward stagewise selection*: Do forward selection, but after every step, modify the remaining predictors such that they are uncorrelated to the selected predictors.
- ▶ etc.



## Important things to keep in mind:

- ▶ The selected model is not guaranteed to be optimal
  - ▶ There are often several equally good models
- ▶ The procedure does not take into account a researcher's knowledge about the predictors
- ▶ Outliers can have a large impact on the procedure
- ▶ Some predictors should be considered together as a group (e.g. dummy indicators for seasons of the year)
- ▶ The coefficients,  $R^2$ , p-values, CI's, etc are all biased/invalid
- ▶ Should not over-interpret the order that the predictors are included
- ▶ Cannot conclude that all variables included are important, or all excluded variables are unimportant



Allows us to use all  $p$  predictors, but will regularize (i.e. shrink) their coefficients in some way.

- ▶ Common to shrink them towards 0



Allows us to use all  $p$  predictors, but will regularize (i.e. shrink) their coefficients in some way.

- ▶ Common to shrink them towards 0

**Question:** Why would shrunk coefficients be better?

- ▶ Will introduce bias, but can significantly reduce the variance
  - ▶ If the variance is noticeably larger, this decreases the test error
- ▶ There are Bayesian motivations to do this: the prior tends to shrink the parameters.



Allows us to use all  $p$  predictors, but will regularize (i.e. shrink) their coefficients in some way.

- ▶ Common to shrink them towards 0

**Question:** Why would shrunk coefficients be better?

- ▶ Will introduce bias, but can significantly reduce the variance
  - ▶ If the variance is noticeably larger, this decreases the test error
- ▶ There are Bayesian motivations to do this: the prior tends to shrink the parameters.

Three common shrinkage methods:

1. Ridge regression
2. Lasso regression
3. Elastic net



Ridge regression solves the following optimization:

$$\min_{\beta} \sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{i,j} \right)^2 + \lambda \sum_{j=1}^p \beta_j^2 \quad (13)$$

In blue: the model RSS

In red: the squared  $\ell_2$  norm of  $\beta$ , or  $\|\beta\|_2^2$

The parameter  $\lambda > 0$  is a tuning parameter. It modulates the importance of fit vs. shrinkage.

- Typically determined via e.g. cross-validation



Writing our loss function in matrix form

$$(\mathbf{Y} - \mathbf{X}\beta)^\top (\mathbf{Y} - \mathbf{X}\beta) + \lambda\beta^\top \beta \quad (14)$$

it can be shown that

$$\hat{\beta}_n^{ridge} = (\mathbf{X}^\top \mathbf{X} + \lambda \mathbf{I}_n)^{-1} \mathbf{X}^\top \mathbf{Y} \quad (15)$$

- ▶ So ridge regression simply adds a positive constant to  $\mathbf{X}^\top \mathbf{X}$ , making it non-singular.





Under the linear model, the mean and covariance of  $\hat{\beta}_n^{ridge}$  are:

$$\begin{aligned}\mathbb{E}[\hat{\beta}_n^{ridge} | \mathbf{X}] &= (\mathbf{X}^\top \mathbf{X} + \lambda \mathbf{I}_n)^{-1} \mathbf{X}^\top \mathbb{E}[\mathbf{Y} | \mathbf{X}] \\ &= (\mathbf{X}^\top \mathbf{X} + \lambda \mathbf{I}_n)^{-1} \mathbf{X}^\top \mathbf{X} \beta\end{aligned}\tag{16}$$

$$\begin{aligned}\text{Cov}[\hat{\beta}_n^{ridge} | \mathbf{X}] &= (\mathbf{X}^\top \mathbf{X} + \lambda \mathbf{I}_n)^{-1} \mathbf{X}^\top \text{Cov}[\mathbf{Y} | \mathbf{X}] \\ &\quad \mathbf{X}^\top (\mathbf{X}^\top \mathbf{X} + \lambda \mathbf{I}_n)^{-1} \\ &= \sigma^2 (\mathbf{X}^\top \mathbf{X} + \lambda \mathbf{I}_n)^{-1} \mathbf{X}^\top \mathbf{X} (\mathbf{X}^\top \mathbf{X} + \lambda \mathbf{I}_n)^{-1}\end{aligned}\tag{17}$$



In least-squares regression, scaling the variables has no effect on the fit of the model:

$$Y = X_0 + \beta_1 X_1 + \beta_2 X_2 + \cdots + \beta_p X_p \quad (18)$$

e.g. **Multiplying  $X_1$  by  $c$  can be compensated by dividing  $\hat{\beta}_1$  by  $c$**

- ▶ i.e. Doing this results in the same RSS



In least-squares regression, scaling the variables has no effect on the fit of the model:

$$Y = X_0 + \beta_1 X_1 + \beta_2 X_2 + \cdots + \beta_p X_p \quad (18)$$

e.g. **Multiplying  $X_1$  by  $c$  can be compensated by dividing  $\hat{\beta}_1$  by  $c$**

- ▶ i.e. Doing this results in the same RSS

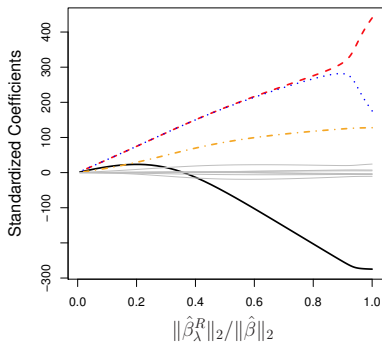
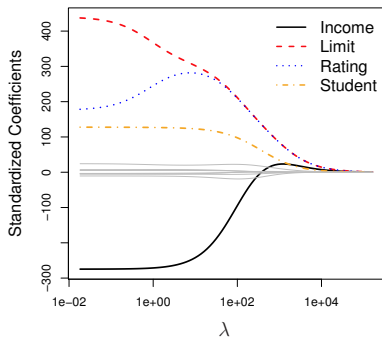
This is not true for ridge regression!

- ▶ Due to  $\|\beta\|_2^2$
- ▶ **In practice:** standardize all predictors (i.e. center and scale such that it has sample variance 1)
  - ▶ e.g. *glmnet* (by Hastie, Tibshirani, and Friedman)

# Example: Ridge regression

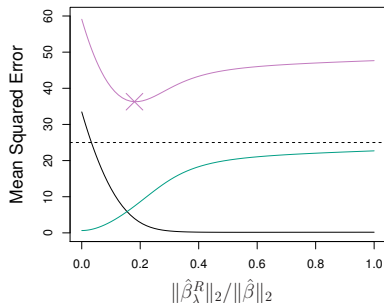
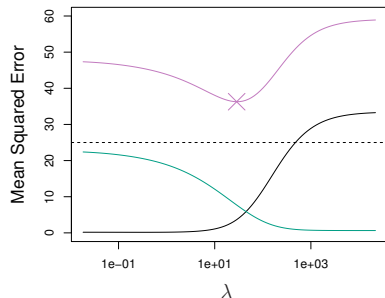


Ridge regression of default in the Credit dataset.



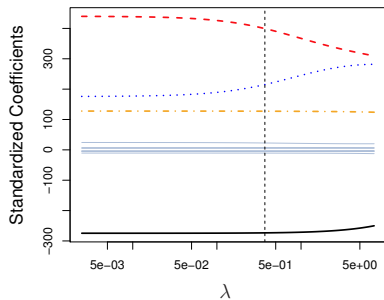
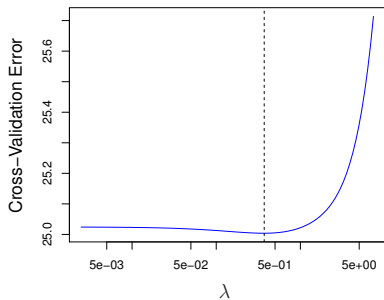


Computing the bias, variance, and test error as a function of  $\lambda$  (in simulation).



Cross validation would yield an estimate of the test error.

# Selecting $\lambda$ by cross-validation





The **L**east **A**bsolute **S**hrinkage and **S**election **O**perator regression solves the following optimization:

$$\min_{\beta} \sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{i,j} \right)^2 + \lambda \sum_{j=1}^p |\beta_j| \quad (19)$$

In blue: the model RSS

In red: the  $\ell_1$  norm of  $\beta$ , or  $\|\beta\|_1$



The **L**east **A**bsolute **S**hrinkage and **S**election **O**perator regression solves the following optimization:

$$\min_{\beta} \sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{i,j} \right)^2 + \lambda \sum_{j=1}^p |\beta_j| \quad (19)$$

In blue: the model RSS

In red: the  $\ell_1$  norm of  $\beta$ , or  $\|\beta\|_1$  **Note:** Unlike ridge regression, LASSO does not have a closed form solution.

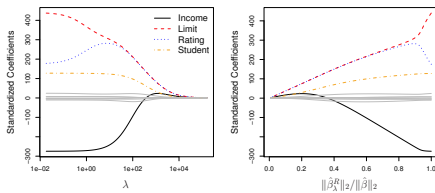
Why would we use the Lasso instead of Ridge regression?

- ▶ Ridge regression shrinks all the coefficients to a non-zero value
- ▶ The Lasso shrinks some of the coefficients all the way to zero.
  - ▶ Similar to subset selection: will select variables for you

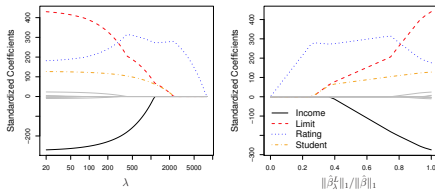




## Ridge regression of default in the Credit dataset.



## Lasso regression of default in the Credit dataset.





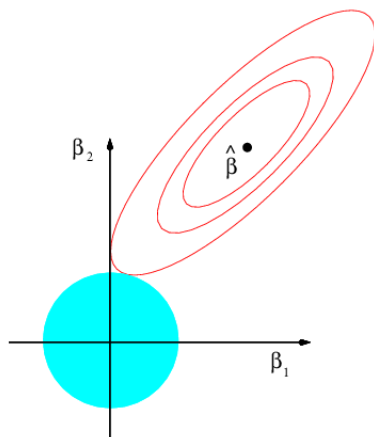
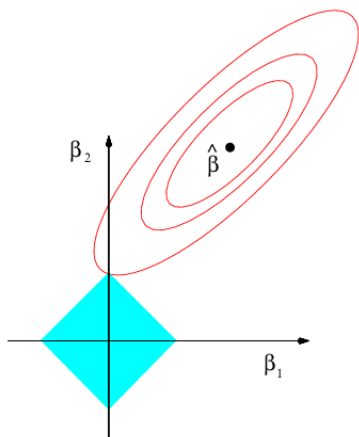
- ▶ **Ridge:** for every  $\lambda$ , there is an  $s$  such that  $\hat{\beta}_\lambda^R$  solves:

$$\min_{\beta} \left\{ \sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{i,j} \right)^2 \right\} \text{ subject to } \sum_{j=1}^p \beta_j^2 < s \quad (20)$$

- ▶ **Lasso:** for every  $\lambda$ , there is an  $s$  such that  $\hat{\beta}_\lambda^L$  solves:

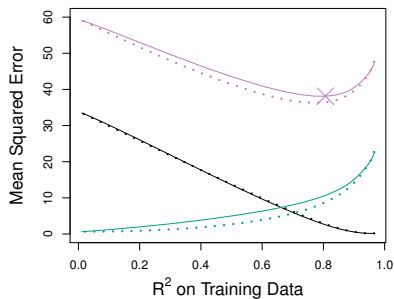
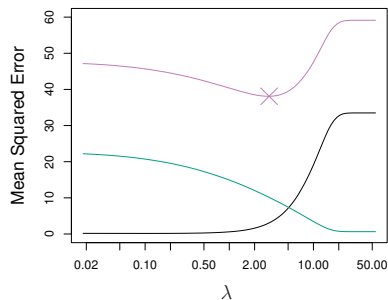
$$\min_{\beta} \left\{ \sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{i,j} \right)^2 \right\} \text{ subject to } \sum_{j=1}^p |\beta_j| < s \quad (21)$$

# An alternative formulation for regularization





**Example 1.** Most of the coefficients are non-zero.

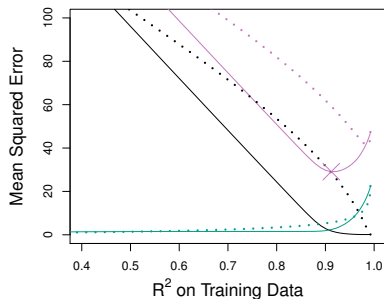
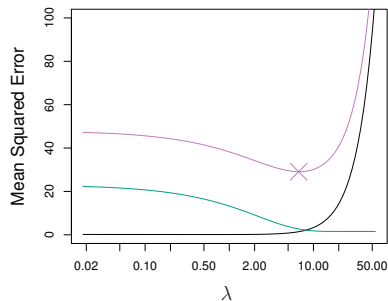


- ▶ Bias, Variance, MSE. The Lasso (—), Ridge (···).
- ▶ The bias is about the same for both methods.
- ▶ The variance of Ridge regression is smaller, so is the MSE.

# When is the Lasso better than Ridge?



**Example 2.** Only 2 coefficients are non-zero.



- ▶ Bias, Variance, MSE. The Lasso (—), Ridge (···).
- ▶ The bias, variance, and MSE are lower for the Lasso.



Combines  $\|\beta\|_2^2$  (ridge) and  $\|\beta\|_1$  (lasso) penalties.

Elastic net solves the following optimization:

$$\min_{\beta} \sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{i,j} \right)^2 + \lambda_2 \sum_{j=1}^p \beta_j^2 + \lambda_1 \sum_{j=1}^p |\beta_j| \quad (22)$$

In blue: the model RSS

In red: both  $\|\beta\|_2^2$  and  $\|\beta\|_1$

This provides a nice trade off between sparsity and grouping.



Combines  $\|\beta\|_2^2$  (ridge) and  $\|\beta\|_1$  (lasso) penalties.

Elastic net solves the following optimization:

$$\min_{\beta} \sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{i,j} \right)^2 + \lambda_2 \sum_{j=1}^p \beta_j^2 + \lambda_1 \sum_{j=1}^p |\beta_j| \quad (22)$$

In blue: the model RSS

In red: both  $\|\beta\|_2^2$  and  $\|\beta\|_1$

This provides a nice trade off between sparsity and grouping.

Typically, we define  $\alpha = \frac{\lambda_2}{\lambda_2 + \lambda_1}$  and instead optimize:

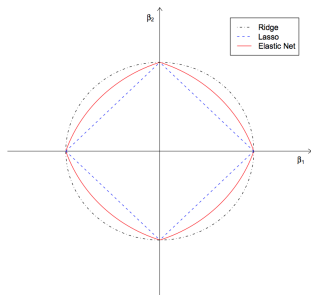
$$\min_{\beta} \sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{i,j} \right)^2 + \alpha \sum_{j=1}^p \beta_j^2 + (1 - \alpha) \sum_{j=1}^p |\beta_j| \quad (23)$$



## Elastic net:

$$\min_{\beta} \left\{ \sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{i,j} \right)^2 \right\} \text{ s.t. } \alpha \|\beta\|_2^2 + (1 - \alpha) \|\beta\|_1 < s \quad (24)$$

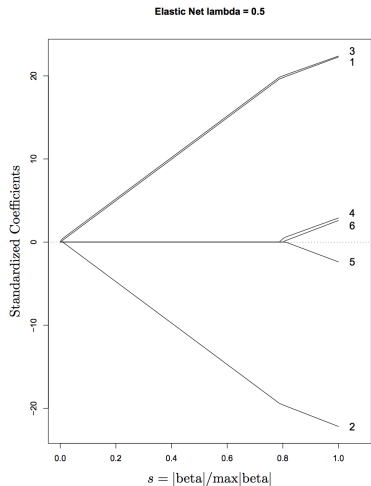
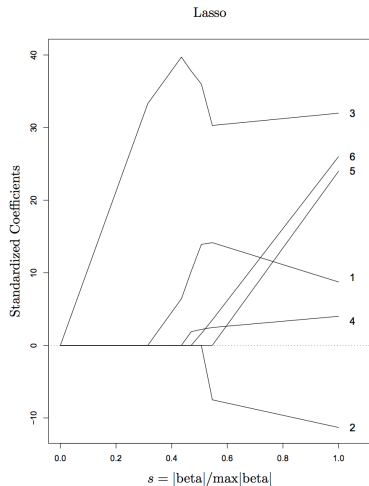
2-dimensional illustration  $\alpha = 0.5$



- ▶ Singularities at the vertexes (to encourage sparsity)
- ▶ Strict convex edges (to encourage grouping)
  - ▶ The strength of convexity varies with  $\alpha$



# Example: Elastic net





$$\min_{\beta} \sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{i,j} \right)^2 + \lambda_2 \sum_{j=1}^p \beta_j^2 + \lambda_1 \sum_{j=1}^p |\beta_j| \quad (25)$$

Method	Shrinkage parameters
OLS	$\lambda_1 = \lambda_2 = 0$
Ridge	$\lambda_1 = 0, \lambda_2 > 0$
LASSO	$\lambda_1 > 0, \lambda_2 = 0$
Elastic net	$\lambda_1 > 0, \lambda_2 > 0$
$\hat{\beta}_n = 0$	$\lambda_1 = \infty$ or $\lambda_2 = \infty$



- ▶ If desired, we could instead consider  $L_q$  penalties for values other than 0, 1, and 2 (e.g.  $q \in (1, 2)$  or  $q > 2$ ).
- ▶ Regularization methods such as the elastic net have been extended to generalized linear models (GLM) as well.
- ▶  $L_1$  and  $L_2$  penalties are also used in contexts other than linear models (e.g. neural networks).
- ▶ As usual, we are faced with the bias-variance tradeoff when choosing our shrinkage parameters,  $\lambda_1$  and  $\lambda_2$ .
- ▶ Other regularized methods are also available, e.g.
  - ▶ Non-negative Garotte Regression
  - ▶ Least Angle Regression
  - ▶ Best subset



*Degrees of freedom* give us a measure of our model's complexity, i.e. the number of free parameters to fit on our data.

- ▶ For OLS, the degrees of freedom is equal to  $p + 1$ .
- ▶ In regularized regression, our parameters are estimated in a restricted manner, controlled by  $\lambda_1$  and  $\lambda_2$ .
  - ▶ Effectively reduced the degrees of freedom in our model
- ▶ We can still compare across models using an *effective degrees of freedom*:

$$df(y, \hat{y}) = \frac{1}{\sigma^2} \sum_{i=1}^n \text{Cov}[y_i, \hat{y}_i | x_i] \quad (26)$$

- ▶ In the case of OLS, this can be shown to reduce to the "standard" degrees of freedom, i.e.  $p + 1$ .



[1] ISL. Chapters 6.

[2] ESL. Chapter 18.